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Online cognitive-based intervention for depression: exploring possible circularity in mechanisms of change

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Background. This study investigates possible circularity in mechanisms of change in participants of Master Your Mood (MYM), a cognitive-based, online intervention for young adults with depressive symptoms. A previous study showed that MYM effectively reduced depression and anxiety and strengthened mastery.

Method. We randomized 244 participants with depressive symptoms into MYM or a wait-list control condition. We explored the circularity hypothesis by several analyses. Correlations were computed to determine the association between (change in) depression and anxiety. Path analysis mediation models were used to explore whether change in anxiety and mastery mediated the intervention effect on depression, whether depression and mastery mediated the effect on anxiety and whether depression and anxiety mediated the effect on mastery. We used linear regression to explore whether early changes in anxiety predicted later changes in depression, and whether early changes in depression predicted later changes in anxiety.

Results. Co-morbidity between depression and anxiety was high (69.2%) and the association between depression and anxiety change was strong ($r=0.677$, $p<0.01$). Changes in anxiety and mastery mediated change in depression (mediation proportion 44%); changes in depression mediated change in anxiety (79%) and mastery (75%). We did not find an early change in anxiety predictive for a late change in depression, and vice versa.

Conclusions. This study appears to confirm the hypothesized circularity in the recovery process. We found high co-morbidity and strong correlation between depression and anxiety levels and bi-directional relationships between potential mediators and outcomes. Early anxiety change scores were not predictive of late depression change scores, and vice versa.

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Key words: Adolescents, cognitive behavioral therapy, depression, mechanisms of change, mediation analysis.

Introduction

Cognitive behavioural therapy (CBT) interventions have been shown in a multitude of studies to be an effective treatment for depression in adults and adolescents (Weisz *et al.* 2006; Calear & Christensen, 2010; Cuijpers *et al.* 2011). However, our knowledge of mechanisms that could explain this recovery is still limited (Kazdin, 2007; Lemmens *et al.* 2011). To evaluate how change comes about, research may

focus on mediators. Kazdin (2007) defines a mediator as an intervening variable that may statistically account for the relationship between the independent and dependent variable. He sees assessing mediation as an important intermediate step between showing a causal relationship and correctly understanding the mechanisms through which the effect occurs. Identifying mediators is important to the further improvement of treatments and their clinical and cost-effectiveness. It enables components to be included that are crucial to recovery (Kazdin, 2007; Lemmens *et al.* 2011). Our present study will focus on the possible mediating role of changes in perceived control, anxiety and depression in the recovery process of participants with symptoms of depression that followed

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the CBT-based online Master Your Mood (MYM) intervention.

Perceived control

In previous mediation studies on CBT interventions perceived control has been explored as a possible mediator (Meulenbeek *et al.* 2010; Warmerdam *et al.* 2010). Perceived control refers to beliefs about one's own ability to control one's environment (Thompson & Schlehofer, 2008). People have a sense of control when they believe that personal action generally controls outcomes (internal locus of control) and that they have the personal skills (self-efficacy) to carry out those actions (Thompson & Schlehofer, 2008). 'Mastery' is a similar concept and is defined as the feeling of the extent to which a person perceives himself to be in control of events and ongoing situations (Moser & Dracup, 1995). As both concepts are often used indiscriminately, no distinction had been made between perceived control and mastery in this study.

Perceived control seems to be a crucial element in understanding anxiety and depression. Low perceived control predicts anxiety and depression symptoms (Chaney *et al.* 1996; Rivard & Cappeliez, 2007). Other research has demonstrated an inter-relationship between stress, perceived control and emotional disturbances (Lachman & Weaver, 1998; Thompson *et al.* 2006; Mausbach *et al.* 2006; 2007; Rosenbaum *et al.* 2012). For example, in the study of Rosenbaum *et al.* (2012) in patients with non-cardiac chest pain, it was hypothesized that perceived control would mediate the relationship between stress and anxiety disorder severity and between stress and mood disorder severity. This study showed that stress and perceived control were both associated with severity of anxiety and mood disorder. It appeared that perceived control fully mediated the relationship between stress and mood disorder severity, but not the severity of anxiety disorder. Furthermore, a few studies were found in which perceived control was examined as a possible mediator in the intervention effect on depression or anxiety. A study on a CBT intervention for depression demonstrated that perceived control mediated the intervention effect on depression (Warmerdam *et al.* 2010). Another study identified perceived control as a mediator in the recovery from anxiety (Bakker *et al.* 2002). A study on panic disorder did not find that feelings of mastery in general mediated the effect of the CBT-based intervention on panic symptomatology; however, perceived self-efficacy in coping with panic did have a mediating effect (Meulenbeek *et al.* 2010).

Our own previous trial on the CBT-based MYM intervention (Van der Zanden *et al.* 2012) suggested that a sense of control may mediate recovery from depression. In that study, the experimental group

included people who did not attend a single session, but who nonetheless showed favourable intervention effects. We surmised that this might be explained by the difference in study conditions to which the experimental group and the wait-list group had been assigned: though both groups attended no sessions, the experimental group made an active decision on it, and thus might have experienced a sense of control, while the wait-listed group did not. Finally, we found studies reporting apparent bi-directional relationships between change on cognitive measures and perceived control and changes on symptom measures for depression and anxiety, indicating that the change process might be circular in nature (Jarrett *et al.* 2007; Anholt *et al.* 2008; Meulenbeek *et al.* 2010).

Summarizing, it seems relevant to investigate the mediating role of perceived control in the intervention effect on depression and anxiety symptoms. In the light of possible circularity in the process of change, it seems also relevant to investigate the reversed condition, that is the mediating role of changes in depression and anxiety symptoms in the intervention effect on perceived control.

Anxiety and depression

Another issue to analyse in building knowledge on mediators and possible circularity in change processes is the strong relationship between depression and anxiety (Axelson & Birmaher, 2001; Hale *et al.* 2009; Hoek *et al.* 2012). As anxiety and depression are strongly associated and processes of change may occur in a circular manner, it is worthwhile to investigate not only perceived control as a mediator for depression and anxiety recovery (and vice versa) as stated before, but also whether change in anxiety mediates depression change and whether change in depression mediates anxiety change. Although we did not find previous mediation studies on this topic, we did discover other research on depression and anxiety that may help in building knowledge on mechanisms of change. Research has shown that a co-morbid anxiety disorder is noted in 25% to 50% of depressed adolescents and 10% to 15% of adolescents with an anxiety disorder have a co-morbid depressive disorder (Brady & Kendall, 1992; Cole *et al.* 1998; Axelson & Birmaher, 2001). Most studies of adolescents have found evidence that anxiety disorder symptoms precede symptoms of depressive disorder (Cole *et al.* 1998), whilst results on the converse relationship are less conclusive (Axelson & Birmaher, 2001). A longitudinal study by Hoek *et al.* (2012) of 497 teenagers drawn from the general adolescent population has found, using non-diagnostic instruments, that symptoms of depression and anxiety developed simultaneously.

As research indicates a strong relationship between depression and anxiety, it has been questioned whether anxiety and depression actually represent two distinct syndromes, or simply the same disorder on a continuum of severity (Lee & Rebok, 2002). More recent research, however, suggests that depression and anxiety do develop as two distinct disorders with parallel growth processes, each with its own unique growth characteristics (Hale *et al.* 2009).

Regarding the strong relationship between depression and anxiety, and possible circularity in processes of change, it seems relevant to explore in the present study the associations between anxiety and depression and the possible mediating roles played by anxiety reduction in depression recovery and vice versa.

Study object and hypotheses

This paper focuses on the mechanisms of change in participants receiving the MYM intervention, an online CBT-based group course for young adults with depressive symptoms. MYM has been shown effective in reducing depressive and anxiety symptoms and increasing mastery (Van der Zanden *et al.* 2012). To explore the hypothesized circularity in the recovery process, we first analyse the relationship between depression and anxiety in our sample. We then perform mediation analyses to detect: (1) any mediating role of changes in anxiety and mastery in changes in depression; (2) any mediating role of changes in depression and mastery in changes in anxiety; and (3) any mediating role of anxiety and depression in changes in mastery. We finally explore whether early changes in anxiety predict later change in depression and whether early changes in depression predict later change in anxiety.

Assuming possible circularity in the recovery process, we expect: (a) high co-morbidity and a strong association between depression and anxiety symptom levels, and changes in them, in our sample; (b) bi-directional changes in mediators and outcomes; and (c) no prediction of late change in depression by early change in anxiety and no prediction of late change in anxiety by early change in depression. The latter expectation is based on the assumption that when relationships are bi-directional and the process of change is circular in nature, it is less likely to find in a period of time that early mediator change is predictive for later change in outcome.

Method

Participants and procedures

A comprehensive description of the participants and procedures can be found in a previous publication

(Van der Zanden *et al.* 2012). Participants were recruited from the general population using promotional materials in general practitioner (GP) offices and educational institutions. Banners and links were placed on mental health websites and on websites popular with youth. Inclusion criteria were: age 16–25 years, informed consent (including parental consent for those aged under 18 years) and a Center for Epidemiologic Studies Depression Scale (CES-D) score of 10–45 (Bouma *et al.* 1995). Applicants were excluded on indications of suicidal ideation with intent and plan, as assessed with the Mini-International Neuropsychiatric Interview (MINI-Plus; Sheehan *et al.* 1998; Van Vliet *et al.* 2000). Assessments took place before randomization (baseline), 12 weeks later (post-intervention) and 12 weeks after that (follow-up). Questionnaires were completed online; the MINI-Plus was administered in online chat sessions. A total of 244 participants were randomized to the MYM course ($n=121$) or a wait-list control group ($n=123$). A flow chart of the sample selection is shown in Fig. 1 (Van der Zanden *et al.* 2012).

Intervention

The online MYM group course is a structured form of CBT for depression. At the core of MYM is the cognitive restructuring of thinking patterns. Course participants are encouraged to detect their own unproductive, unrealistic thoughts, and they are then taught to transform these into realistic, helpful thoughts. Performance of pleasant daily activities is also encouraged, and a mood measure is completed daily to clarify the connection between pleasant activities and mood. The course we evaluated took place at fixed times in a secured chatroom, which participants entered with usernames and passwords. Anonymity was ensured by self-chosen nicknames. The course comprised six 90-min sessions, at set times every week, and home exercises. The sessions were structured around six themes (Van der Zanden *et al.* 2012). During the sessions, course material was introduced by the facilitators and displayed in the chatroom using text and images. Participants could respond, share experiences and ask questions. Emoticons could be used to express feelings. Participants and professionals could read session transcripts afterwards. The course was guided by one or two trained professionals, depending on group size (six participants maximum).

Outcome measures

Depressive symptoms

Symptoms of depression were assessed with the 20-item CES-D (Radloff, 1977; Bouma *et al.* 1995),

Fig. 1. Flow chart of respondent selection. MYM, Master Your Mood; CES-D, Center for Epidemiologic Studies Depression Scale; MINI-Plus, Mini-International Neuropsychiatric Interview.

which measures the frequency of 20 depressive symptoms over the past week using four-point Likert scales. Total score range is 0–60, with higher ratings indicating more depressive symptoms. Computerized and paper CES-D versions correlate at a very high level (Ogles *et al.* 1998). The web-based CES-D has proved a reliable, valid screening instrument in a Dutch adolescent population (Cuijpers *et al.* 2008).

Anxiety symptoms

The anxiety subscale of the Hospital Anxiety and Depression Scale (HADS-A; Zigmond & Snaith, 1983) was used to assess anxiety symptoms. The Dutch version of HADS has been validated (Spinhoven *et al.* 1997). The anxiety subscale consists of seven items measuring anxiety symptoms on a four-point Likert scale, with a 0–21 score range, and with higher ratings indicating higher states of anxiety.

Perceived control

The Dutch version of the five-item Mastery Scale was used to assess perceived control (Pearlin & Schooler, 1978). Mastery refers to beliefs about one's own ability to control one's environment. Responses are rated on a five-point Likert scale, with a total score range of 5–25, and higher scores indicating a greater sense of mastery. The Mastery Scale has good psychometric properties (Pearlin & Schooler, 1978).

Statistical analysis

All analyses were based on an intention-to-treat sample (including data from all participants, whether or not they received the intervention). The expectation-maximization method was used to impute data missing at post-treatment and follow-up (Van der Zanden *et al.* 2012). It imputes values by maximum-likelihood estimation using the observed data in an iterative process (Dempster *et al.* 1977). The 12-week assessment was used as the post-treatment measurement and the 24-week assessment as the follow-up. The follow-up was performed in the experimental group only, as the control group had access to the course after post-treatment. Baseline between-group differences in demographics and potential mediators were investigated using χ^2 and t tests. Correlation analysis was used to investigate associations between levels of anxiety and depression at baseline and post-treatment. For the calculation of the co-morbidity rate of depression with anxiety, we used the cut-off scores CES-D ≥ 22 (Cuijpers *et al.* 2008) and HADS ≥ 11 (Crawford *et al.* 2001).

Mediation analysis was then conducted, to investigate the mechanisms through which the treatment affects the outcome variable(s), following the steps for testing multiple mediators outlined by Baron & Kenny (1986):

- (1) Treatment condition should predict change in the outcome variable.

- (2) Treatment condition should predict change in the possible mediators.
- (3) Change in possible mediators should predict change in the outcome variable.
- (4) The effect of treatment condition on change in outcome should be attenuated when the effect of mediators is statistically controlled.

Path analysis mediation models (MacKinnon, 2008) were used to explore the relationship between the outcome variable and two potential correlated mediators. To investigate possible circularity in the recovery process, in our study each of the variables depression, anxiety and mastery was tested alternately as an outcome variable and as a potential mediator. In the first analysis, the outcome was the difference between baseline and post-treatment scores for depression, and potential mediators were baseline to post-treatment differences in anxiety and mastery scores. In the second analysis, the outcome was anxiety differences, with depression and mastery differences as possible mediators. In the third analysis, the outcome was differences in mastery and potential mediators were differences in depression and anxiety.

Univariate analyses were performed with SPSS 19 (SPSS Inc., USA) and mediation analyses with Mplus (Muthén and Muthén, USA). A two-sided significance level of $p=0.05$ was used. Mediation proportions for each mediator were computed by dividing the absolute value of the indirect effect, $ABS(a_1*b_1)$, by the total effect, $ABS(c)+ABS(a_1*b_1)+ABS(a_2*b_2)$ (Bate *et al.* 2009). In the mediation models, the a-path refers to the effect of the independent variable on the mediator, the b-path represents the effect of the mediator on the dependent variable after controlling for the effect of the independent variable, and the c-path refers to the direct effect of the independent variable on the dependent variable after controlling for mediators (see Figs 2–4). Variance inflation factors (VIFs) were calculated to assess possible collinearity effect in multiple mediation analysis, meaning that the indirect effect attenuates to the extent that the mediators are correlated in the model (Belsey *et al.* 1980; Preacher & Hayes, 2008).

Results

Baseline participant characteristics

Demographic and clinical characteristics of experimental and control group members are shown in Table 1 (Van der Zanden *et al.* 2012). At baseline, t tests for independent groups and χ^2 tests revealed no significant between-group differences on demographic and outcome variables (CES-D, HADS-A, mastery).

Co-morbidity and association between depression and anxiety

We first analysed co-morbidity – defined as the presence of one or more disorders in addition to a primary disease or disorder – of depression with the additional disorder anxiety. For clinical depression we used a CES-D cut-off score of ≥ 22 (Cuijpers *et al.* 2008) and for anxiety disorder a HADS cut-off score of ≥ 11 (Crawford *et al.* 2001). Although the inclusion criterion was set on CES-D >10 , 86.5% (211/244) of the participants had a CES-D score ≥ 22 . Furthermore, 85.2% (208/244) had a HADS score ≥ 11 .

Then we calculated the co-morbidity rate of depression with anxiety, according to the given definition. At baseline we found a co-morbidity rate of 69.2% (146/211), meaning that of the group with a CES-D score ≥ 22 ($n=211$), 146 respondents had a HADS score ≥ 11 .

Next, we computed the association between depression and anxiety scores at baseline and at post-treatment, finding strong correlations between depression and anxiety at both assessments ($r=0.60$, $p<0.001$; $r=0.71$, $p<0.001$). Finally, we computed correlations between the baseline to post-treatment change scores in depression and anxiety. The correlation between the changes was strong ($r=0.68$, $p<0.01$) and in the expected direction: greater changes in depression corresponded with greater changes in anxiety.

Association between treatment and outcomes

As previously reported (Van der Zanden *et al.* 2012), MYM participants showed significantly greater improvements at 12 weeks than controls, with a large between-group effect size for depressive symptoms [CES-D, $d=0.94$, 95% confidence interval (CI) 0.64–1.23] and moderate effect sizes for anxiety (HADS-A, $d=0.49$, 95% CI 0.24–0.75) and mastery ($d=0.44$, 95% CI 0.19–0.70). Improvements in the MYM group were maintained at 24 weeks. In the present study, the outcomes depression, anxiety and mastery were also considered as possible mediators, which means that the same estimated treatment effect applies to the association between treatment and mediators as well (Table 2).

Associations between treatment and mediating variables

The next step in examining the hypothesized bi-directional relationships between change in outcomes and change in mediators was to analyse the various associations between treatment, potential mediators and outcomes. The first and second requirements for mediation analysis of Baron & Kenny (1986) were checked, and had been fulfilled (Table 2). The check of the third and fourth requirements was performed

Table 1. Baseline characteristics of the 244 participants

	Experimental (<i>n</i> =121)	Control (<i>n</i> =123)	All (<i>n</i> =244)	Statistics
Female gender	101 (83.5)	105 (85.4)	206 (84.4)	$\chi^2_1=0.17, p=0.68$
Mean age, years (s.d.)	20.8 (2.2)	21.0 (2.3)	20.9 (2.2)	$t_{242}=0.64, p=0.53$
Age groups				
16–17 years	5 (4.1)	4 (3.3)	9 (3.7)	$\chi^2_2=0.16, p=0.92$
18–21 years	66 (54.5)	69 (56.1)	135 (55.3)	
22–25 years	50 (41.3)	50 (40.7)	100 (41.0)	
Education level ^a				$\chi^2_2=0.63, p=0.73$
Low	10 (8.2)	10 (8.1)	20 (8.1)	
Middle	50 (41.3)	45 (36.6)	95 (38.9)	
High	61 (50.4)	68 (55.3)	129 (52.8)	
Daily activities				$\chi^2_2=1.90, p=0.38$
Study	83 (68.6)	85 (69.1)	168 (68.9)	
Paid employment	32 (26.4)	27 (22.0)	59 (24.2)	
Other	6 (4.9)	11 (8.9)	17 (7.0)	
Living situation				$\chi^2_3=3.54, p=0.32$
With parents	56 (46.3)	59 (48.0)	115 (47.1)	
With partner	13 (10.7)	18 (14.6)	31 (12.7)	
Alone	26 (21.5)	16 (13.0)	42 (17.2)	
With others	26 (21.5)	30 (24.4)	56 (23.0)	
Experienced in web-chatting	63 (52.1)	64 (52.0)	127 (52.0)	$\chi^2_1=0.00, p=0.99$
Prior professional help	75 (62.0)	76 (61.8)	151 (61.9)	$\chi^2_1=0.00, p=0.98$
Current professional help at baseline	36 (29.8)	39 (31.7)	77 (31.6)	$\chi^2_1=0.11, p=0.74$
Mean CES-D depression score (s.d.) ^b	32.5 (8.4)	32.3 (8.2)	32.3 (8.3)	$t_{242}=0.28, p=0.77$
Mean HADS-Anxiety (s.d.) ^c	11.2 (3.6)	11.8 (3.7)	11.5 (3.6)	$t_{242}=1.27, p=0.21$
Mean mastery (s.d.) ^d	12.8 (3.4)	12.8 (3.6)	12.8 (3.5)	$t_{242}=0.17, p=0.86$

Data are given as number (percentage), unless otherwise indicated.

s.d., Standard deviation.

^a Highest completed or present education: low=primary or lower secondary school or less; middle=secondary school or intermediate vocational school; high=professional school or university.

^b Dutch version of the 20-item Center for Epidemiologic Studies Depression Scale (Bouma *et al.* 1995).

^c Dutch version of the seven-item Hospital Anxiety and Depression Scale (Spinhoven *et al.* 1997).

^d Five-item Mastery Scale (Pearlin & Schooler, 1978).

in the same equation. We studied three possible mediation models (Figs 2–4). Table 3 summarizes the indirect effects (or mediation effects) and the mediation proportions.

Model 1: depression as mediated by anxiety and mastery

As Fig. 2 shows, the association between treatment and change in depression decreases from -7.804 to -4.69 after control for anxiety and mastery, but remains significant ($p<0.001$). Table 3 reveals a significant indirect effect of anxiety (-2.11 , 95% CI -03.64 to -0.80); the mediation proportion is 0.25, meaning that 25% of the effect of treatment on depression is explained by the change in anxiety. The indirect effect of mastery

is also significant (-1.63 , 95% CI -2.93 to -0.48), with a mediation proportion of 19%. Altogether, then, the mediators anxiety and mastery explain 44% of the intervention effect on depression.

Model 2: anxiety as mediated by depression and mastery

As Fig. 3 shows, the association between treatment and anxiety is no longer significant after control for depression and mastery, with the regression coefficient decreasing from -2.051 to 0.28 ($p=0.532$). Table 3 shows the mediating effects of depression and mastery. Only the depression effect is significant ($p<0.001$). The mediation proportions are 0.79 for depression and 0.08 for mastery. Altogether, the

Table 2. Estimated treatment effect on outcome variables and estimated treatment effect on potential mediators from baseline to post-measurement

Variable	Group $\times$ time		Between-group effect size	
	Estimate	<i>t</i>	<i>p</i>	Difd (95% CI)
Depression: MYM <i>v.</i> WL	-7.804	-7.184	<0.001	0.94 (0.64–1.23)
Anxiety: MYM <i>v.</i> WL	-2.051	-5.070	<0.001	0.49 (0.24–0.75)
Mastery: MYM <i>v.</i> WL	1.438	3.508	0.001	0.44 (0.19–0.70)

MYM, Master Your Mood course; WL, wait-list control group; Difd, effect size differences between the MYM group and the wait-list control group; CI, confidence interval.

Fig. 2. Graphic display of the estimated path analysis model. The two regression equations, with depression as outcome and anxiety and mastery as potential mediators, are represented by single-headed arrows. Parameter estimates, standard errors (and *p* values) are reported for each regression equation.

mediators depression and mastery explain 87% of the intervention effect on anxiety. The VIF for depression and anxiety was 1.79, meaning that there is no collinearity effect.

Model 3: mastery as mediated by anxiety and depression

Fig. 4 shows that the association between treatment and change in mastery is no longer significant after control for anxiety and depression. The coefficient decreases from 1.44 to -0.31 ($p=0.512$). Table 3 reveals a significant indirect effect of depression (1.61, 95% CI 0.91–2.46), with a mediation proportion of 0.75, and a non-significant indirect effect of anxiety

Fig. 3. Graphic display of the estimated path analysis model. The two regression equations, with anxiety as outcome and depression and mastery as potential mediators, are represented by single-headed arrows. Parameter estimates, standard errors (and *p* values) are reported for each regression equation.

Fig. 4. Graphic display of the estimated path analysis model. The two regression equations, with mastery as outcome and anxiety and depression as potential mediators, are represented by single-headed arrows. Parameter estimates, standard errors (and *p* values) are reported for each regression equation.

(0.22, 95% CI -0.02 to 0.57), with a mediation proportion of 0.10. Altogether, the mediators anxiety and depression thus explain 85% of the intervention effect on mastery. The VIF for depression and mastery was 1.65, meaning that there is no collinearity effect.

Relationship between early changes in anxiety or depression and late changes in depression or anxiety

To detect whether an early change in anxiety was predictive of a late change in depression, we first calculated correlations of baseline to post-treatment change

Table 3. Intervention effects on depression, anxiety and mastery as mediated by changes in depression, anxiety and mastery, and indirect effects of the intervention on outcomes through proposed mediators (a*b paths)

Mediators	Depression outcome				Anxiety outcome				Mastery outcome			
	Estimate	(95% CI)	p ^a	Prop	Estimate	(95% CI)	p ^a	Prop	Estimate	(95% CI)	p ^a	Prop
Depression	–	–	–	–	–1.81	(–2.65 to –1.10)	0.000	0.79	1.61	(0.91 to 2.46)	0.000	0.75
Anxiety	2.11	(3.64–0.80)	0.002	0.25	–	–	–	–	0.22	(–0.02 to 0.57)	0.151	0.10
Mastery	1.63	(0.63–2.93)	0.009	0.19	0.192	(–0.01 to 0.48)	0.15	0.08	–	–	–	–
Both	3.74	(1.65–5.88)	0.000	0.44	2.00	(1.24 to 2.92)	0.000	0.87	1.83	(1.12 to 2.70)	0.000	0.85

CI, Confidence interval; Prop, mediation proportion.

^a A p value of 0.000 means that the p value is <0.001.

scores on these measures. The correlation was strong ($r = 0.68$, $p < 0.01$) and in the expected direction. We then calculated correlations between ‘early’ change in anxiety from baseline to post-treatment and ‘late’ change in depression between post-treatment and follow-up. The correlation was significant but weak ($r = -0.231$, $p < 0.01$) and not in the expected direction, indicating that a greater reduction in anxiety scores during the first 12 weeks corresponded to a smaller reduction in depression scores in the following 12 weeks. We next computed the reverse relationship: whether early change in depression predicted later change in anxiety. Again, the correlation was significant but weak ($r = -0.288$, $p < 0.001$) and in the unexpected direction: a greater early decrease in depression scores corresponded with a smaller late decrease in anxiety scores. As early changes in anxiety or depression were negatively related to late changes in the other condition, no additional regression analyses were performed.

Discussion

Main results

The purpose of this study was to help clarify mechanisms of change in online CBT for depression. We analysed the process of change in participants in the MYM intervention, an online CBT-based group course for adolescents with symptoms of depression. A randomized controlled trial of MYM had already shown a significant reduction in depressive and anxiety symptoms in the intervention participants and an increase in their sense of control, or mastery (Van der Zanden *et al.* 2012). To gain further insights into the process of change, which we expected to be circular in nature, we analysed the relationship between depression and anxiety in our sample and conducted mediation analyses using three different models. These explored possible mediators and outcomes and their inter-relationships. We also analysed whether changes in the mediator anxiety preceded changes in the outcome depression, and vice versa.

The co-morbidity of depression with anxiety was high and changes in depression and anxiety were strongly intercorrelated. The results of the multiple mediation analyses indicated that changes in anxiety and mastery mediated change in depression (mediation proportion 44%); changes in depression mediated anxiety change (79%) and mastery change (75%). Mastery change appeared not to mediate significant change in anxiety and vice versa (8% and 10%). We did not find that early changes in anxiety or depression scores predicted late changes in the other condition.

Comparison with other research

This study appears to confirm the hypothesized circularity in the recovery process in online CBT interventions. First, we detected a strong intercorrelation between the changes in depression and anxiety occurring in our sample. Although our participants had been recruited and selected exclusively on depressive symptoms and not on anxiety, we found high co-morbidity between the two, at a baseline rate of 69.2%, exceeding that reported for the general youth population (Axelson & Birmaher, 2001). However, it seems in line with a study of Birmaher *et al.* (1996) on clinical youth samples, in which a co-morbidity rate up to 70% was reported for depression with anxiety. The MYM trial participants seemed to fit well into the group of at-risk adolescents with the highest symptom severities who are most vulnerable to increasing levels of the other disorder, as described by Hale *et al.* (2009). It should be noticed that the latter and current study were based on self-report screening instruments (which may generate false-positive diagnosis) and that Axelson & Birmaher (2001) and Birmaher *et al.* (1996) based their results on (diagnostic) interviews. Nevertheless, the co-morbidity rate in our study seems to be substantial, which makes it worthwhile when recruiting adolescents for CBT interventions like MYM not to focus exclusively on those with depressive symptoms, but on those with anxiety as well.

A second indication of circularity in the recovery process is that the three mediating models revealed bi-directional relationships between changes in mediators and outcomes, consistent with other studies (Jarrett *et al.* 2007; Anholt *et al.* 2008; Meulenbeek *et al.* 2010). Depression was the mediator that exhibited the largest mediation proportion in the three models, explaining 79% of the intervention effect on anxiety. This suggests that anxiety reduction is explained largely by improvement in depression. Change in anxiety explained 25% of the intervention's effect on depression.

The result that the mediating impact of anxiety change on depression change was less than the reverse condition might have been affected by the fact that the MYM intervention was originally designed to reduce depression symptoms and that the effect size for anxiety was less than that for depression (see Table 2). Nevertheless, the results of the mediation analyses suggest that targeting depressive symptoms can be a way of alleviating anxiety symptoms.

Another finding was that change in mastery appeared to explain a larger proportion of the effect on depression (19%) than of that on anxiety (8%). This is roughly in line with other studies: the

mediating role of mastery in the change in depression seems to be more consistent (Warmerdam *et al.* 2010; Rosenbaum *et al.* 2012) than the mediating role of mastery in anxiety change (Meulenbeek *et al.* 2010; Rosenbaum, *et al.* 2012). This might be explained by the scales for assessing perceived control used in our and the other studies, which seems to be focused predominantly on mastery of depression (sense of control) and not on mastery as it is required to master anxiety. While mastery of anxiety does require a sense of control, it also requires the ability to experience anxiety-provoking situations without avoiding them (Meulenbeek *et al.* 2010). In the MYM intervention, the last aspect is targeted by stimulating participants to enhance pleasant activities in daily life, which may include anxiety-provoking activities that were found pleasant before participants' distress emerged.

A third indication of circularity is our finding that early changes in anxiety or depression did not predict later changes in the other disorder. This was consistent with findings by Warmerdam *et al.* (2010) and with our expectation that when relationships are bi-directional and the change process is circular in nature, it is less likely to find in a time schedule one variable is predictive for the other. The result could, however, also be due to the fact that most of the change in our sample took place between baseline and post-treatment, leaving little change to detect from post-treatment to follow-up.

Limitations

In addition to the limitations described in our previous publication on the randomized controlled trial of MYM (Van der Zanden *et al.* 2012), the current study may have several other limitations. First, the absence of assessments during the course of the intervention prevented us from analysing the precise sequence of changes. Our study contained three assessments: baseline, post-intervention and follow-up. Additional interim measurements could have given more detailed insights into the process of recovery from the beginning (when most changes appear to occur; Garratt *et al.* 2007; Warmerdam *et al.* 2010) to the end of the intervention and follow-up. A more rigorous test of mediation would require that changes in specific mediatory variables temporally precede changes in the outcome variables. The best design for mediation analysis would entail a fine-grained analysis of changes in mediators and symptoms, based on several interim measurements. This would allay concerns that both the mediator and the symptoms changed simultaneously or due to a third variable (Kazdin, 2007).

Second, the study population was selected on depressive symptoms, and not on levels of anxiety or mastery. Although a majority of the participants (69.2%) had co-morbid anxiety, the inclusion criteria could have still affected the outcomes of the three mediation models. Whereas respondents with low scores on depression (CES-D < 10) were excluded from the trial, those with low anxiety scores were not. For the latter, there was not much room or need to reduce their anxiety symptoms. This could have led to the smaller mediating impact of the anxiety change score on depression as compared with the other direction. Otherwise, respondents with CES-D scores > 45 were excluded, but not those with high HADS-A scores.

Third, only a limited set of potential mediators was available in this study. The three mediating models containing the variables mastery, anxiety and depression were set up to analyse statistical connections between these variables, and thereby to enhance knowledge about mediation and mechanisms of change. The high co-morbidity between depression and anxiety and the substantial effects of the intervention on both disorders may prompt more active recruitment of people with anxiety, in addition to those with depression. Nevertheless, no clues for concrete improvements in the intervention's components emerged. Therefore, future research with more potential mediators is necessary.

Future directions

In the field of psychological treatment, knowledge of mediators and mechanisms of change is still very limited. 'Next-step' research is recommended to help unfold such mechanisms. As Kazdin (2007, p. 6) has emphasized, 'the case for a mediator is built by a sequence of studies that may vary in the set of criteria they address and the clarity of the demonstration'. He underlined that several studies are needed, meeting a range of criteria (strong associations, specificity, consistency, experimental manipulation, gradient, demonstrated timelines and plausibility), before one can conclude that an intervening process explains change.

In research on CBT interventions like MYM, such 'next-step' research might include more frequent assessments of symptoms and potential mediators, in order to build knowledge on the sequence of change. Special points of interest would be how to prevent attrition caused by the more frequent assessments during the intervention and what influence the assessments themselves might have on the intervention effect.

Since specific mediators like cognitions and perceived control may not account for the full effect in

CBT interventions (Longmore & Worrell, 2006; Warmerdam *et al.* 2010), future research should include non-specific variables as well. These could help uncover underlying mechanisms (Stephen & Craighead, 2006). Previous studies (Spek *et al.* 2007; Van der Zanden *et al.* 2012) have implied that social support from professionals or course participants could be a relevant mediator. Perceived stress, which could itself be affected by feelings of support, might also mediate the intervention effect (Mausbach *et al.* 2007). Feelings of hope and expectation (Stephen & Craighead, 1995) or a feeling of control derived from the mere act of applying for the intervention (even if the applicant ultimately does not attend any session; Van der Zanden *et al.* 2012) could further explain intervention effects.

The strong connection between anxiety and depression, though these appear to be different disorders (Hale *et al.* 2009), suggests underlying dimensions involving factors such as the psychobiological response to threat and stress. Repeated dysregulation of those systems may result in arousal and emotional reactions like depression (Dubovsky, 1990; Cole *et al.* 1998). Future research should therefore incorporate physiological outcomes into the total concept of change mechanisms, in addition to the psychological outcomes discussed so far. Interdisciplinary research will therefore be of interest in future work.

Given the strong association between depression and anxiety, a final recommendation would be to develop and evaluate treatment approaches that target both conditions simultaneously.

Conclusions

This study appears to confirm the hypothesized circularity in the recovery process.

Depression change scores corresponded strongly to the change scores of co-morbid anxiety. Furthermore, depression change scores mediated significantly anxiety and mastery change scores, anxiety change scores mediated significantly depression change scores and mastery change scores mediated significantly depression change scores.

Because specific mediators such as cognitions and sense of control do not appear to explain the entire effect of CBT interventions, and because underlying psychological and physiological mechanisms may play a role in the recovery process, future research should include non-specific and physiological variables. A more profound understanding of mediators will further the refinement of theoretical notions relating to the complex mechanisms of change. This will help to optimize treatment and improve its (cost) effectiveness.

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Declaration of Interest

R.v.d.Z. is one of the developers of the online MYM group course, and the Trimbos Institute is a cooperation partner in MYM, but none of these derives financial income from the MYM intervention. There are no competing interests.

References

- Anholt GE, Kempe P, de Haan E, van Oppen P, Cath DC, Smit JH, van Balkom AJLM (2008). Cognitive versus behavior therapy: processes of change in the treatment of obsessive-compulsive disorder. *Psychotherapy and Psychosomatics* 77, 38–42.
- Axelsson DA, Birmaher B (2001). Relation between anxiety and depressive disorders in childhood and adolescence. *Depression and Anxiety* 14, 67–78.
- Bakker A, Spinhoven P, Van der Does AJW, Van Balkom AJLM, Van Dyck R (2002). Locus of control orientation in panic disorder and the differential effects of treatment. *Psychotherapy and Psychosomatics* 71, 85–89.
- Baron RM, Kenny DA (1986). The moderator–mediator variable distinction in social psychological research: conceptual, strategic, and statistical considerations. *Journal of Personality and Social Psychology* 51, 1173–1182.
- Bate SL, Stigler MH, Thompson MS, Arora M, Perry CL, Reddy KS, MacKinnon DP (2009). Psychosocial mediators of a school-based tobacco prevention program in India: results from the first year of Project MYTRI. *Prevention Science* 10, 116–128.
- Belsey DA, Kuh E, Welsch RE (1980). *Regression Diagnostics: Identifying Influential Data and Sources of Collinearity*. John Wiley and Sons: New York.
- Birmaher B, Ryan ND, Williamson D, Brent D, Kaufman J, Dahl R, Perel J, Nelson B (1996). Childhood and adolescent depression: a review of the past 10 years – part I. *Journal of the American Academy of Child and Adolescent Psychiatry* 35, 1427–1439.
- Bouma J, Ranchor AV, Sanderman R, van Sonderen E (1995). *Het meten van symptomen van depressie met de CES-D. Een handleiding (Measurement of Symptoms of Depression With the CES-D: A Manual)*. Rijksuniversiteit Groningen: Groningen.
- Brady EU, Kendall PC (1992). Comorbidity of anxiety and depression in children and adolescents. *Psychological Bulletin* 111, 244–255.
- Calear AL, Christensen H (2010). Review of Internet-based prevention and treatment programs for anxiety and depression in children and adolescents. *Medical Journal of Australia* 192 (Suppl.), S12–S14.
- Chaney JM, Mullins LL, Uretsky DL, Doppler MJ, Palmer WR, Wees SJ (1996). Attributional style and depression in rheumatoid arthritis: the moderating role of perceived illness control. *Rehabilitation Psychology* 41, 205–223.
- Cole DA, Peeke LG, Martin JM, Truglio R, Seroczynski AD (1998). A longitudinal look at the relation between depression and anxiety in children and adolescents. *Journal of Consulting and Clinical Psychology* 66, 451–460.
- Crawford JR, Henry JD, Crombie C, Taylor EP (2001). Brief report normative data for the HADS from a large non-clinical sample. *British Journal of Clinical Psychology* 40, 429–434.
- Cuijpers P, Boluijt P, Van Straten A (2008). Screening of depression in adolescents through the Internet: sensitivity and specificity of two screening questionnaires. *European Child and Adolescent Psychiatry* 17, 32–38.
- Cuijpers P, van Straten A, Donker T, Andersson G (2011). Psychological treatment of depression: results of a series of meta-analyses. *Nordic Journal of Psychiatry* 65, 354–364.
- Dempster AP, Laird NM, Rubin DB (1977). Maximum likelihood from incomplete data via the EM algorithm. *Journal of the Royal Statistical Society Series B (Statistical Methodology)* 39, 1–38.
- Dubovsky SL (1990). Understanding and treating depression in anxious patients. *Journal of Clinical Psychiatry* 51 (Suppl.), 3–83–8, 14–17.
- Garratt G, Ingram RE, Rand KL, Sawalani G (2007). Cognitive processes in cognitive therapy: evaluation of the mechanisms of change in the treatment of depression. *Clinical Psychology Science and Practice* 14, 224–239.
- Hale W, Raaijmakers Q, Muris P, Van Hoof A, Meeus W (2009). One factor or two parallel processes? Comorbidity and development of adolescent anxiety and depressive disorder symptoms. *Journal of Child Psychology and Psychiatry* 50, 1218–1226.
- Hoek W, Van Lier PAC, Koot HM (2012). Combined trajectories of depressive and anxiety symptoms from early to middle adolescence: descriptions, correlates and gender differences. In *Subclinical Depression and Anxiety in Adolescence: Developmental Trajectories and Online Intervention*, Dissertation (ed. W. Hoek), pp. 25–50. Faculty of Psychology and Education, Vrije Universiteit: Amsterdam.
- Jarrett RB, Vittengl JR, Doyle K, Clark LA (2007). Changes in cognitive content during and following cognitive therapy for recurrent depression: substantial and enduring, but not predictive of change in depressive symptoms. *Journal of Consulting and Clinical Psychology* 75, 432–446.
- Kazdin AE (2007). Mediators and mechanisms of change in psychotherapy research. *Annual Review of Clinical Psychology* 3, 1–27.

- Lachman ME, Weaver SL** (1998). The sense of control as a moderator of social class differences in health and well-being. *Journal of Personality and Social Psychology* **74**, 763–773.
- Lee LC, Rebok GW** (2002). Anxiety and depression in children: a test of the positive–negative affect model. *Journal of the American Academy of Child and Adolescent Psychiatry* **41**, 419–426.
- Lemmens LHJM, Arntz A, Peeters FPML, Hollon SD, Roefs A, Huibers MJH** (2011). Effectiveness, relapse prevention and mechanisms of change of cognitive therapy vs. interpersonal therapy for depression: study protocol for a randomised controlled trial. *Trials* **12**, 150.
- Longmore RJ, Worrell M** (2007). Do we need to challenge thoughts in cognitive behavior therapy? *Clinical Psychology Review* **27**, 173–187.
- MacKinnon DP** (2008). *Introduction to Statistical Mediation Analysis*, chapter 6. Lawrence Erlbaum Associates: New York.
- Mausbach BT, Mills PJ, Patterson TL, Aschbacher K, Dimsdale JE, Ancoli-Israel S, von Känel R, Grant I** (2007). Stress-related reduction in personal mastery is associated with reduced immune cell β 2-adrenergic receptor sensitivity. *International Psychogeriatrics* **19**, 935–946.
- Mausbach BT, Patterson TL, Von Känel R, Mills PJ, Dimsdale JE, Ancoli-Israel S** (2006). Personal mastery attenuates the effects of caregiving stress on psychiatric morbidity. *Journal of Nervous and Mental Disease* **194**, 132–134.
- Meulenbeek P, Spinhoven P, Smit F, van Balkom A, Cuijpers P** (2010). Cognitive mediation of panic reduction during an early intervention for panic. *Acta Psychiatrica Scandinavica* **122**, 20–29.
- Moser DK, Dracup K** (1995). Psychosocial recovery from a cardiac event; the influence of perceived control. *Heart and Lung* **24**, 273–280.
- Ogles BM, France CR, Lunnen KM, Bell MT, Goldfarb M** (1998). Computerized depression screening and awareness. *Community Mental Health Journal* **34**, 27–38.
- Pearlin LI, Schooler C** (1978). The structure of coping. *Journal of Health and Social Behavior* **19**, 2–21.
- Preacher KJ, Hayes AF** (2008). Asymptotic and resampling strategies for assessing and comparing indirect effects in multiple mediator models. *Behavior Research Methods* **40**, 879–891.
- Radloff LS** (1977). The CES-D Scale: a self-report depression scale for research in the general population. *Applied Psychological Measurement* **1**, 385–401.
- Rivard V, Cappeliez P** (2007). Perceived control and coping in women faced with activity restriction due to osteoarthritis: relations to anxious and depressive symptoms. *Canadian Journal of Aging* **26**, 241–253.
- Rosenbaum DL, White KS, Gervino EV** (2012). The impact of perceived stress and perceived control on anxiety and mood disorders in noncardiac chest pain. *Journal of Health Psychology* **17**, 1183–1192.
- Sheehan DV, Lecrubier Y, Sheehan KH, Amorim P, Janavs J, Weiller E** (1998). The Mini-International Neuropsychiatric Interview (M.I.N.I.): the development and validation of a structured diagnostic psychiatric interview for DSM-IV and ICD-10. *Journal of Clinical Psychiatry* **59**, 22–33.
- Spek V, Cuijpers P, Nyklíček I, Riper H, Keyzer J, Pop V** (2007). Internet-based cognitive behaviour therapy for symptoms of depression and anxiety: a meta-analysis. *Psychological Medicine* **37**, 319–328.
- Spinhoven P, Ormel J, Sloekers PP, Kempen GI, Speckens AE, Van Hemert AM** (1997). A validation study of the Hospital Anxiety and Depression Scale (HADS) in different groups of Dutch subjects. *Psychological Medicine* **27**, 363–370.
- Stephen SI, Craighead WE** (1995). The role of nonspecific factors in cognitive–behavior therapy for depression. *Clinical Psychology: Science and Practice* **1**, 138–155.
- Stephen SI, Craighead WE** (2006). Rapid early response, cognitive modification, and nonspecific factors in cognitive behavior therapy for depression: a reply to Tang and DeRubeis. *Clinical Psychology: Science and Practice* **6**, 295–299.
- Thompson SC, Schlehofer MM** (2008). The many sides of control motivation. Motives for high, low and illusory control. In *Handbook of Motivation Science* (ed. J. Y. Shah and W. L. Gardner), pp. 41–56. Guilford Press: New York.
- Thompson SC, Schlehofer MM, Bovin MJ, Dougan BT, Montes D, Trifskin S** (2006). Disposition, control strategies, and distress in the general public after the 2001 terrorist attack. *Anxiety, Stress and Coping* **19**, 143–159.
- Van der Zanden AP, Kramer J, Gerrits R, Cuijpers P** (2012). Effectiveness of an online group course for depression in adolescents and young adults: a randomized trial. *Journal of Medical Internet Research* **14**, e86.
- Van Vliet IM, Leroy H, van Megen H** (2000). *M.I.N.I. Plus. International Neuropsychological Interview. Nederlandse Versie 5.0.0. (Dutch Version 5.0.0)*. UMC: Utrecht.
- Warmerdam L, van Straten A, Jongsma J, Twisk R, Cuijpers P** (2010). Online cognitive behavioral therapy and problem-solving therapy for depressive symptoms: exploring mechanisms of change. *Journal of Behavior Therapy and Experimental Psychiatry* **41**, 64–70.
- Weisz JR, McCarty CA, Valeri SM** (2006). Effects of psychotherapy for depression in children and adolescents: a meta-analysis. *Psychological Bulletin* **132**, 132–149.
- Zigmond AS, Snaith RP** (1983). The Hospital Anxiety and Depression Scale. *Acta Psychiatrica Scandinavica* **67**, 361–370.